



Modeling the potential health benefits of lower household air pollution after a hypothetical liquified petroleum gas (LPG) cookstove intervention[☆]



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ABSTRACT

Introduction: Improved biomass and advanced fuel cookstoves can lower household air pollution (HAP), but levels of fine particulate matter (PM_{2.5}) often remain above the World Health Organization (WHO) recommended interim target of 35 µg/m³.

Methods: Based on existing literature, we first estimate a range of likely levels of personal PM_{2.5} before and after a liquified petroleum gas (LPG) intervention. Using simulations reflecting uncertainty in both the exposure estimates and exposure-response coefficients, we estimate corresponding expected health benefits for systolic blood pressure (SBP) in adults, birthweight, and pneumonia incidence among children < 2 years old. We also estimate potential avoided premature mortality among those exposed.

Results: Our best estimate is that an LPG stove intervention would decrease personal PM_{2.5} exposure from approximately 270 µg/m³ to approximately 70 µg/m³, due to likely continued use of traditional open-fire stoves. We estimate that this decrease would lead to a 5.5 mmHg lower SBP among women over age 50, a 338 g higher birthweight, and a 37% lower incidence of severe childhood pneumonia. We estimate that decreased SBP, if sustained, would result in a 5%–10% decrease in mortality for women over age 50. We estimate that higher birthweight would reduce infant mortality by 4 to 11 deaths per 1000 births; for comparison, the current global infant mortality rate is 32/1000 live births. Reduced exposure is estimated to prevent approximately 29 cases of severe pneumonia per year per 1000 children under 2, avoiding approximately 2–3 deaths/1000 per year. However, there are large uncertainties around all these estimates due to uncertainty in both exposure estimates and in exposure-response coefficients; all health effect estimates include the null value of no benefit.

Conclusions: An LPG stove intervention, while not likely to lower exposure to the WHO interim target level, is still likely to offer important health benefits.

1. Introduction

Observational studies suggest potential health benefits from lower levels of household air pollution (HAP) associated with clean cookstoves; the strongest evidence is for lung cancer, chronic obstructive pulmonary disease (COPD), and childhood pneumonia (Bruce et al., 2015). However, the reduction in exposure with improved biomass stoves has not been in general sufficient to reduce levels to the WHO annual interim fine particulate matter (PM_{2.5}) target (IT-1) level of 35 µg/m³ (Bruce et al., 2015; Clark et al., 2013a; Balakrishnan et al., 2014; Pope et al., 2017). In a 2014 literature review, Rehfuess et al.

(2014) estimated that the impact of an intervention with improved biomass cookstoves with chimneys vs. traditional (open-fire) stoves reduces personal PM_{2.5} exposure among women from 270 µg/m³ to 120 µg/m³, based on five studies with data on personal exposure. Possible reasons for less-than-desired reductions in exposure include poor stove performance (design and maintenance), stove stacking (use of traditional stoves along with new stoves), and possible other sources of PM_{2.5} beyond cooking (Clark et al., 2013a). Increasing evidence from combustion studies have raised questions about whether any stoves that rely on biomass (wood, dung, agricultural waste, coal, charcoal) can achieve targeted reductions in PM_{2.5} without other interventions such

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as improved ventilation (Yip et al., 2017). This less-than-desired performance of improved biomass cookstoves has motivated recent trials of stoves using non-biomass fuels, such as liquid petroleum gas (LPG).

There have been relatively few randomized intervention trials of cookstoves that have measured health effects. Smith et al., 2011 focused on childhood pneumonia in Guatemala, Thompson et al. (2011) focused on birth outcomes in Guatemala, Smith-Sivertsen et al. (2009) focused on respiratory function in Guatemala, McCracken et al. (2007) focused on blood pressure in Guatemala, Mortimer et al. (2017) focused on childhood pneumonia in Malawi, and Romieu et al. (2009) focused on adult lung function in Mexico. There is also a recently completed trial in Nigeria (Alexander et al., 2017). While the Mortimer et al. trial did not observe a protective effect for pneumonia, the other trials have shown some health benefits of improved stoves. Perhaps the best known evidence of protective effects is for childhood pneumonia in the RESPIRE trial (Smith et al., 2011). There have also been beneficial effects on blood pressure in older women (McCracken et al., 2007) and in pregnant women (Alexander et al., 2017), beneficial effects on birthweight (Thompson et al., 2011), and mixed results on lung function (Romieu et al., 2009; Smith-Sivertsen et al., 2009).

The WHO target level for an annual average PM_{2.5} exposure is 10 µg/m³, but given the difficulties in achieving this target level, WHO has set an annual interim target (IT-1) level of 35 µg/m³ (WHO, 2014). Because of the steep slopes of estimated exposure-response curves at low levels of PM_{2.5} for many health outcomes, combined with a plateau of the curve at higher levels (Burnett et al., 2014), Bruce et al. (2015) have argued that substantial health gains are likely to be observed only if personal exposure levels fall below the WHO interim target level of 35 µg/m³, particularly for childhood pneumonia, ischemic heart disease (IHD), and stroke. However, Bruce et al. (2015) also noted that interventions resulting in lower levels of exposure (from 200 to 300 µg/m³ to 35–80 µg/m³, without necessarily attaining the 35 µg/m³ target level), nonetheless can be expected to result in 20% to 50% decreases in risk of several outcomes (COPD, lung cancer, infant pneumonia, birth outcomes). This conclusion is consistent with the WHO *Guidelines for Indoor Air Quality* (WHO, 2014), which also notes that lowering personal exposures from 200 to 300 µg/m³ to 35–80 µg/m³ is likely to result in reductions in risk of 20–50% for several outcomes (WHO, 2014).

Liquefied petroleum gas (LPG) interventions are likely to result in lower exposures than have been observed with interventions with improved biomass cookstoves (Balakrishnan et al., 2014; Rehfuess et al., 2014). Absent any additional exposure from traditional stoves (stove stacking) for cooking or heating, and absent any important ambient air contribution (e.g., trash burning, traffic), it's likely that LPG stove would lead to indoor levels of PM_{2.5} at or near the WHO target of 10 µg/m³. Even with stove stacking and elevated ambient levels, we believe that use of LPG stoves will result in substantially reduced concentrations compared to traditional or improved biomass stoves. Consequently, it is likely that LPG stove interventions can show appreciable health benefits for several outcomes, with more expected improvement than from improved biomass stoves. Interventions with LPG are becoming more feasible as the price of LPG has dropped considerably worldwide in the past 5 years (BP Global, 2017). Two trials using LPG as an intervention have recently been completed, although results have not been published (Tielsch et al., 2014; Jack et al., 2015), and there is an ongoing trial using an ethanol-burning stove (HAP, 2017). The large, multi-centric trial focused exclusively on cooking with LPG stoves – the four-country HAPIN (Household Air Pollution Intervention) study – has just begun (HAPIN, 2017). The present authors are part of the team conducting that study.

Randomized trials are typically short-term and cannot easily evaluate long-term effects on chronic disease like lung cancer, stroke, and heart disease. However, such trials can show benefits for outcomes that can be measured in the short term, such as childhood pneumonia, birth outcomes, and adult blood pressure. These short-term outcomes can

either show immediate direct benefits, such as pneumonia and low birth weight (both of which can affect infant mortality risk), or potential longer term benefits (blood pressure, which if lowered consistently results in lower rates of chronic disease).

There are relatively few studies of HAP and health effects that provide exposure-response data based on measured personal exposures (Quansah et al., 2017). However, there are exposure-response data for the short-term outcomes mentioned above – childhood pneumonia, low birth weight, and blood pressure – based on personal exposure. In addition, Burnett et al. (2014) have developed integrated exposure-response curves (IERS) for PM_{2.5} in relation to chronic diseases (lung cancer, stroke, COPD, heart disease) as well as infant pneumonia, by using PM_{2.5} data from studies of HAP, mainstream tobacco smoke, second-hand smoke, and ambient air pollution.

Using published exposure-response data, we estimate the health benefits gained with exposure reductions anticipated to result from LPG interventions, for three outcomes that can be measured in a short-term trial: blood pressure in adult women (> age 50), birth weight, and pneumonia incidence among young children.

2. Methods

2.1. Review of the literature

Our estimates of the effects of HAP on health are drawn from published literature. First, we examined the studies included in several recent systematic reviews, including Balakrishnan et al. (2014), Rehfuess et al. (2014), Pope et al. (2017), Quansah et al. (2017), Bruce et al. (2013), and Bruce et al. (2015). We updated the results of these reviews via a PubMed search using keywords 'PM_{2.5}', 'biomass', 'air pollution', 'cookstoves', 'household air pollution', 'blood pressure', 'birthweight', 'premature birth', 'pre-term birth', 'gas stove' and 'LPG' and 'respiratory infection' and 'pneumonia'. Finally, we queried investigators in the field regarding their knowledge of studies relevant to this investigation.

2.2. Personal PM_{2.5} data

We focus on PM_{2.5} rather than carbon monoxide (CO), or other exposure metrics, because the bulk of the literature on health effects in relation to measured air pollution (ambient and household) refers to PM_{2.5}. We also focus on personal rather than area (kitchen) exposures, as personal exposures are most relevant when considering health effects.

Rehfuess et al. (2014) reviewed the literature for cookstove intervention studies as part of a WHO document. Data based on five studies with either 24 or 48 h sampling indicates that an intervention with improved biomass cookstoves with chimneys would lower personal PM_{2.5} exposure among women from 270 µg/m³ to about 120 µg/m³, a 56% reduction (Fig. 6, Rehfuess et al., 2014). The reduction in personal exposure due to the interventions in Rehfuess et al. (2014) of 60%–64% corresponded well with the reduction observed in kitchen concentration.

Balakrishnan et al. (2014) conducted an analogous WHO review of the cookstove literature based on studies which were not interventions. These authors estimated personal exposure levels from traditional stoves of 267 µg/m³ of PM_{2.5} based on 8 studies, very similar to the estimate of 270 µg/m³ for traditional stoves in the intervention studies reviewed by Rehfuess et al. (2014).

The potential effect of an intervention with LPG stoves on personal PM_{2.5} exposure is not known. The only LPG intervention study cited by Rehfuess et al. (2014) was conducted in Sudan, and measured 24-h kitchen, rather than personal, concentrations (Bates et al., 2017, see Annex 18). Furthermore, the data presented are for all respirable PM, and are not directly translatable to PM_{2.5}. There is also one non-intervention study with data on exposure from LPG stoves cited by

Balakrishnan et al. (2014) (Fig. 9), where the personal exposure of women to $PM_{2.5}$ was approximately $70 \mu\text{g}/\text{m}^3$.

For our best estimate of personal exposure from LPG stoves, we take the estimate of $70 \mu\text{g}/\text{m}^3$ from Balakrishnan et al. (2014).

The recent articles by Pope et al. (2017) and Quansah et al. (2017) reviewed much of the same literature as Balakrishnan et al. (2014) and Rehfuess et al. (2014), but revealed only one additional study (Hartinger et al., 2013) with personal $PM_{2.5}$ exposure over a 48-h period. This study found somewhat lower personal $PM_{2.5}$ levels from traditional stoves (mean $145 \mu\text{g}/\text{m}^3$, $n = 28$). We also found one more recent report from China (Hu et al., 2014), not included in the reviews cited here, which had higher levels of personal $PM_{2.5}$ exposures for unvented stoves/firepits using wood or plant-based biomass (mean $406 \mu\text{g}/\text{m}^3$, geometric mean $335 \mu\text{g}/\text{m}^3$), based on 17 households. On the other hand, kitchen exposures were lower compared to personal exposures (30%–40%), unlike virtually all of the other literature. These differences from the other literature may reflect that in China there were significant sources outside the home.

Given the relatively small numbers of samples in the studies by Hartinger et al. (2013) and Hu et al. (2014), and the unusual findings in the study by Hu et al. (2014), we have chosen to rely on Rehfuess et al.'s estimate of $270 \mu\text{g}/\text{m}^3$ as our best estimate of personal $PM_{2.5}$ exposure from traditional stoves.

We focus on arithmetic rather than geometric means, because these are generally what are reported in the literature. Many papers do not provide the geometric mean, and hence we don't have a good estimate of what the literature overall indicates for the geometric mean for different types of stoves (e.g., Balakrishnan et al., 2014, Table 3). The arithmetic mean, while influenced more by outliers, is comprehensible to policy makers and others who may not be familiar with geometric means, and some standards and guidelines are set in terms of annual arithmetic averages.

2.3. Blood pressure

Exposure-response data for blood pressure (BP) are taken from a panel study by Baumgartner et al. (2011). There were 280 women enrolled in this study (mean age 52), with multiple personal $PM_{2.5}$ measurements and BP measurements. Ever-smokers and pregnant women were excluded. The geometric mean of 24-h personal $PM_{2.5}$ was $55 \mu\text{g}/\text{m}^3$ in the summer and $117 \mu\text{g}/\text{m}^3$ in the winter. Using mixed models the authors found an increase in systolic blood pressure (SBP) of 2.2 (95% CI 0.8–3.7) mmHg for each unit of log $PM_{2.5}$ among all women, and an increase of 4.1 mmHg (95% CI 1.5–6.6) for each unit of log $PM_{2.5}$ for women over 50. There was little effect on SBP in women aged 25–50.

There are no other comparable published studies of non-pregnant women with personal exposure to $PM_{2.5}$ from HAP and with exposure-response data for BP. However, there are several supporting studies. McCracken et al. (2007) studied 120 women of age > 38 years, comparing SBP among those exposed after an intervention, with controls. Personal $PM_{2.5}$ averaged $264 \mu\text{g}/\text{m}^3$ in the control group, and $102 \mu\text{g}/\text{m}^3$ in the intervention group. SBP was 3.7 mmHg lower (95% CI –0.6–8.1) in the intervention arm, and diastolic BP was 3.0 mmHg lower (95% CI 0.4–5.7). The results from Baumgartner et al. (2011) are also consistent with results for women over the age of 40 years in Clark et al. (2013b), which compared SBP before and after an intervention with an improved cookstove (without a control group) among 74 Nicaraguan non-pregnant women (27 over the age of 40 years). Area $PM_{2.5}$ (48-h) was reduced from a geometric mean of $1172 \mu\text{g}/\text{m}^3$ to $208 \mu\text{g}/\text{m}^3$ in a subset of women ($n = 25$) with HAP data (personal $PM_{2.5}$ not measured). SBP dropped an average of 5.9 mmHg (95% confidence interval 0.4–11.3) in women age 40 years and older, after the intervention, with little effect on women under that age. Finally, similar observations were reported in a 2016 abstract regarding (non-pregnant) women and SBP in relation to HAP by Young et al. (<https://epiresearch.org/wp-content/uploads/2016/07/Abstract-Book-Final-070416.pdf>).

These authors studied 104 women cross-sectionally, of whom 35 were older than 40 years. Among the women 40 years and older, SBP was 4.6 (–0.7, 9.9) mmHg higher per increased unit of log $PM_{2.5}$, a finding remarkably similar to Baumgartner et al. (2011) (personal communication Jennifer Peel, March 12, 2017).

It should be noted that there are two studies of HAP exposure and pregnant women (Quinn et al., 2016), Alexander et al., 2017). Both studies found associations between higher exposure and higher blood pressure. These studies are not directly relevant for non-pregnant women, as blood pressure is known to vary during pregnancy, falling initially and rising in the third trimester (Hermida et al., 2000).

Other supporting data come from studies of ambient air pollution and blood pressure. Liang et al. (2014) conducted a meta-analysis of 25 cross-sectional and panel studies and found an increase of 1.39 (0.87–1.91) mmHg per $10 \mu\text{g}/\text{m}^3$ increase in $PM_{2.5}$. Most of these studies used $PM_{2.5}$ levels measured at the time of or slightly before BP measurements. A more recent large study from China (Lin et al., 2017) found a similar increase of 1.33 (0.04–3.56) mmHg for SBP, using an estimated prior 3-year average of $PM_{2.5}$ levels. These studies show a similar effect as Baumgartner et al. (2011) in the low-dose region. For example, Baumgartner et al. (2014) would estimate a 4.6 mmHg increase in SBP going from 10 to $40 \mu\text{g}/\text{m}^3$, while using an average of the results from Liang et al. (2014) and Lin et al. (2017), would result in an estimated 4 mmHg increase in SBP.

Here we focus on the effects of HAP on SBP among older women. Outside of pregnancy, older women are in general those most likely to experience environmental effects on BP, given that BP changes little with age until age 30, when it increases steadily (Pinto, 2007). We therefore use the data from older women (over age 50) in Baumgartner et al. (2011) for our calculations. Given the similarity of the data from ambient air pollution studies in the low dose region, and the lack of relevant data from ambient air pollution studies at higher levels, we have not combined Baumgartner et al. (2014) with those results. From Baumgartner et al. (2011), we used as the estimated non-exposed blood pressure, the lowest SBP of approximately 118 mmHg, as per Fig. 1 in that publication. We calculated expected SBP assuming a 4.1 mmHg increase (1.5–6.6) for each increase in one unit of log $PM_{2.5}$.

2.4. Birthweight

Wylie et al. (2017) measured personal $PM_{2.5}$ and birthweight among 239 pregnant women in Tanzania, of whom 118 had personal $PM_{2.5}$ measurements. The geometric mean $PM_{2.5}$ level was $41 \mu\text{g}/\text{m}^3$ (geo. std.

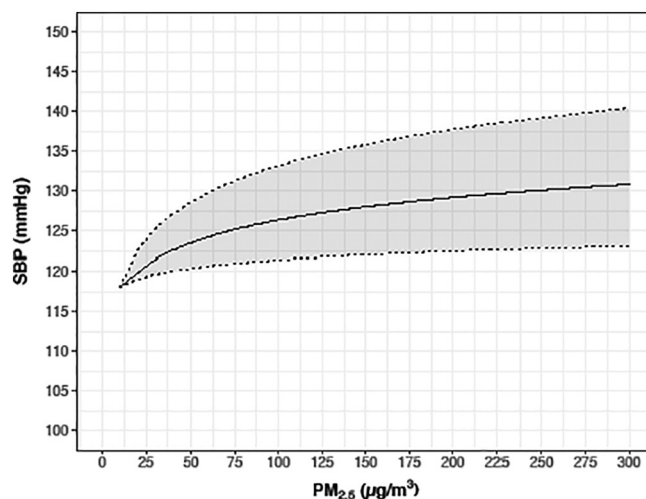


Fig. 1. Exposure-response for systolic blood pressure (SBP) in women > 50 from Baumgartner et al. (2011).

dev. $21 \mu\text{g}/\text{m}^3$), and 87% of measurements exceeded $25 \mu\text{g}/\text{m}^3$. In linear regression using the log of $\text{PM}_{2.5}$, there was a decrease in birth weight of 270 g (95% CI 0–540) for each increase in 1 unit log $\text{PM}_{2.5}$. Results were similar when the data were restricted to full term births.

Supportive evidence comes from a meta-analysis of ambient air and birthweight studies. Sun et al. (2016) found a decrease of 15.9 g (95% CI 5.0, 26.8) birthweight for each increase of $10 \mu\text{g}/\text{m}^3$ $\text{PM}_{2.5}$, across 17 studies, with $\text{PM}_{2.5}$ levels averaged over the entire gestation.

Further supportive evidence comes from a study in Poland of indoor air $\text{PM}_{2.5}$ in relation to birthweight (Jedrychowski et al., 2004). These authors studied 362 pregnant non-smoking women in Poland, with 48-hour personal measurements of $\text{PM}_{2.5}$. $\text{PM}_{2.5}$ concentrations averaged $42 \mu\text{g}/\text{m}^3$, with a range of 10 to 147, and for every increase in log $\text{PM}_{2.5}$ there was a decrease of 200 g (95% CI 15–385) in birthweight, a finding similar to Wylie et al. (2017).

Thompson et al. (2011) also provide some data on personal $\text{PM}_{2.5}$ exposure and birthweight, in a study using a biomass-burning chimney stove; however, there are no exposure-response data beyond the estimated effect of two different exposures. The intervention group had ppm CO levels about 1.5 ppm lower than controls, which would be about 1.69 lower CO measured as mg/m^3 . Using the conversion from Northcross et al., 2010, this is equivalent to an increase of $213 \mu\text{g}/\text{m}^3$ $\text{PM}_{2.5}$. The intervention group had children who weighed 89 g more than the control group, or 328 g more in those women who had the intervention stove in the 1st or 2nd trimester. This would mean an increase in weight of about 15 g per $10 \mu\text{g}/\text{m}^3$ decrease in $\text{PM}_{2.5}$ for women exposed in the 1st or 2nd trimester, comparable to the effect found in Sun et al. (2016), assuming a linear exposure-response (as did Sun et al.).

Other information comes from a meta-analysis of 10 HAP studies from Amegah et al. (2014), who estimated that burning biomass fuel in unimproved stoves, compared to improved stoves or using other cleaner fuel, led to a decrease in birth weight on the average of 86 g (95% CI 55–117 g). However, no exposure levels are provided in this meta-analysis.

We have chosen to use a weighted average of the estimates of effects from Wylie et al. (2017), Sun et al. (2016), and Jedrychowski et al. (2004), to estimate the effect of $\text{PM}_{2.5}$ on birthweight. All of these studies provide exposure-response results based on personal exposure. Two studies used log $\text{PM}_{2.5}$, while one used untransformed $\text{PM}_{2.5}$. For birthweight estimation considering a fixed exposure, we used the inverse of the coefficient of variation (standard deviation/mean) as our weight. We then derived the standard deviation of the weighted average, assuming independence of the three studies, and used that to construct 95% confidence intervals. We used the same weights in simulations considering exposure as a random variable (see below).

2.5. Childhood pneumonia

The exposure-response curve for acute lower respiratory infection (ALRI) (childhood pneumonia, < 18 months) is available from Burnett et al. (2014), based on data from the RESPIRE trial as well as studies of second-hand smoke and ambient air (although the RESPIRE data provide the bulk of the data). Results are from an excess relative risk model. Supplemental detailed relative risk (RR) data from Burnett et al. (2014), for each unit of exposure from 1 to $300 \mu\text{g}/\text{m}^3$, are available (Global Health Data Exchange, 2017). The exposure data in RESPIRE were based on personal measurement of the child's carbon monoxide exposure, which was highly correlated with $\text{PM}_{2.5}$ data. $\text{PM}_{2.5}$ data were estimated from the CO data via a conversion equation; $\text{PM}_{2.5}$ had a 0.87 correlation with CO (R-square 0.76) in these data (Northcross et al., 2010).

It should be noted that the recent randomized trial in Malawi (Mortimer et al., 2017) did not observe a reduction in childhood pneumonia with improved cookstoves, and in fact found borderline-significant higher severe pneumonia in the intervention arm; however,

to date these authors have not presented exposure data, so that the impact of the intervention on exposure levels is not known. Other HAP observational study data do support the relationship observed in the RESPIRE study (e.g., Ezzati and Kammen, 2001), but do not provide personal $\text{PM}_{2.5}$ data.

2.6. Simulations

To incorporate uncertainty in both the exposure-response function and the exposure estimates at specific levels of interest, we used Monte Carlo simulations by drawing realizations for the exposure-response parameters and the exposures of interest from normal distributions. For the exposure-response parameters for birthweight and blood pressure, we assumed they were normally distributed with mean and standard deviation as reported in the literature. For the personal exposure level, the data in both Balakrishnan et al. (2014) and Rehfuess et al. (2014) indicate that in most exposure studies the standard deviation is approximately equal to the mean, and hence we drew from normal distributions under these assumptions. We focused on personal exposures at $270 \mu\text{g}/\text{m}^3$, $120 \mu\text{g}/\text{m}^3$, $70 \mu\text{g}/\text{m}^3$, and $35 \mu\text{g}/\text{m}^3$, and $10 \mu\text{g}/\text{m}^3$ corresponding to plausible levels for unimproved stoves, improved stoves, gas stoves, the WHO interim target level, and the WHO recommended level (the referent). About 15% of simulated exposure levels were below zero and discarded; hence in effect the exposures were drawn from a left-truncated normal distribution at zero. We conducted 10,000 simulations for blood pressure and birthweight. Simulations for pneumonia, in contrast, were based on the publicly available data from Burnett et al. (2014) (Global Health Exchange 2017), where 1000 Monte Carlo simulations of the three parameters in the IER model are provided. For each of these 1000 sets of parameter values, we drew an exposure as noted above, and then calculated the corresponding relative risk. We report the median, the 5th percentile, and the 95th percentile from our simulations for all three outcomes.

3. Results

3.1. Exposure-response estimates

Figs. 1–3 show the estimated exposure-response curves for our three outcomes, based on data in the published literature.

Table 1 provides simulation results for different exposure using the exposure-response coefficients from the literature, for selected exposure levels for all three outcomes, including our best estimates of traditional stove exposure ($270 \mu\text{g}/\text{m}^3$), of an LPG intervention ($70 \mu\text{g}/\text{m}^3$), as well

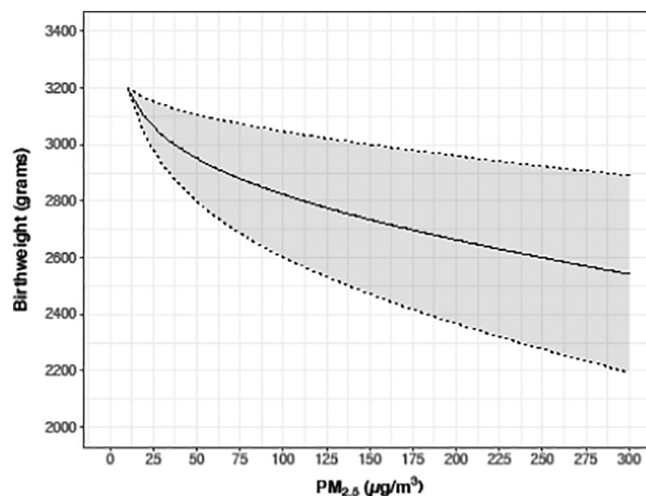


Fig. 2. Exposure-response for birthweight, weighted average of Wylie et al. (2017), Sun et al. (2016), and Jedrychowski et al. (2004).

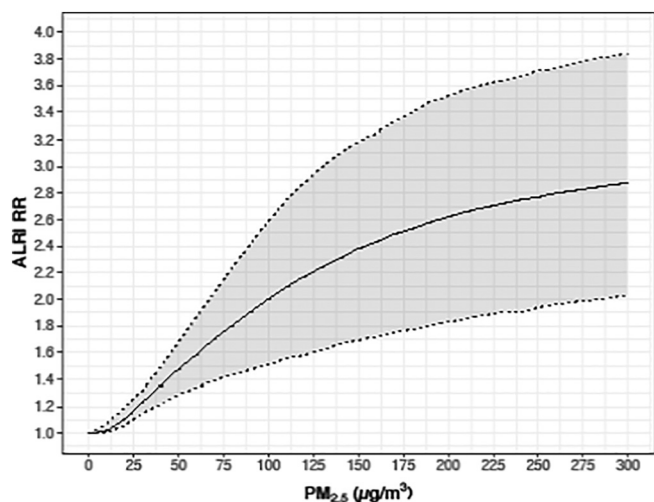


Fig. 3. Exposure-response for childhood pneumonia from Burnett et al. (2014).

as 120 $\mu\text{g}/\text{m}^3$ (improved biomass cookstove). In Supplemental Table 1 we also provide results assuming a fixed exposure level for every 10 unit increment from 10 to 300 $\mu\text{g}/\text{m}^3$.

As a sensitivity analysis for the birthweight data, we excluded the data from the meta-analysis of Sun et al. (2016), on the grounds that the data come from a meta-analysis of ambient air studies rather than indoor air studies, and were based on a model using untransformed $\text{PM}_{2.5}$, rather than $\log \text{PM}_{2.5}$, used in the other two studies. The simulation results indicated somewhat lower median birthweights at each exposure level than when Sun et al. (2016) was included (Table 1). For example, results for 270 $\mu\text{g}/\text{m}^3$ were a median birthweight of 2448 (5th% 1832, 95th% 2993), and for 70 $\mu\text{g}/\text{m}^3$ were 2756 (5th% 2304, 95th% 3186). But the estimated benefit of a decrease of exposure from 270 $\mu\text{g}/\text{m}^3$ to 70 $\mu\text{g}/\text{m}^3$ did not differ greatly. With Sun et al. in the analysis, the benefit was an increase of 338 g, while without Sun et al. the increase was 308 g.

3.2. Potential impacts on health

3.2.1. Blood pressure

Current evidence supports a causal link between $\text{PM}_{2.5}$ and cardiovascular disease and stroke (Hoek et al., 2013), which in turn are also known to be related to higher blood pressure (James et al., 2014). $\text{PM}_{2.5}$ effects on cardiovascular and cerebrovascular disease might be mediated by changes in blood pressure. To date, however, there is limited evidence of an effect of HAP on cardiovascular disease (Fatmi and Coggon, 2016; Mitter et al., 2016; Weichenthal et al., 2017).

Based on our simulations, estimated $\text{PM}_{2.5}$ exposures from biomass stoves (270 $\mu\text{g}/\text{m}^3$) and LPG stoves (70 $\mu\text{g}/\text{m}^3$) would result in estimated SBP among women > 50 years of age of 131.4 and 126.0 mm Hg, respectively (median difference 5.5, 2.5th% – 8.6, 97.5th% 20.8). While the estimated decrease in SBP associated with

lower levels of $\text{PM}_{2.5}$ may appear modest, on a population level (assuming such reductions were sustainable) it would be expected to result in important estimated reductions in risks for ischemic heart disease and stroke. However, the 95% simulation interval is wide and includes the null value.

A recent randomized trial of blood pressure medication among older US adults at 102 clinics across the US and Puerto Rico resulted in a decrease in all-cause mortality of about 1.82% per mm Hg reduction in SBP, so that gas stoves compared to traditional stoves might – in the long term and assuming the point estimate of a reduction of 5.5 is accurate – result in a 10% reduction in mortality (SPRINT group, 2015). Longitudinal analyses of NHANES data suggest an approximate 10% increase in overall mortality risk for every 10 mm Hg increase in SBP for those over age 50, which would suggest about a 5.5% reduction in mortality for users of LPG stoves vs. traditional stoves over age 50 years (Taylor et al., 2011). In comparison, the IER curve for ischemic heart disease death from Burnett et al. (2014), indicates that for 70 year olds the RR for heart disease death would drop approximately 9% with a change in exposure from 270 $\mu\text{g}/\text{m}^3$ to 70 $\mu\text{g}/\text{m}^3$, a finding not dissimilar from the above estimates which were based on extrapolating from blood pressure to heart disease mortality.

It should be noted that the benefits of lowered SBP are likely to be seen even in populations with relative low SBP at baseline. In the SPRINT study referenced above, the protective effect of treatment was strongest in the treatment group which had the lowest BP at baseline (< 132 mm Hg, RR = 0.70).

Increased heart disease is likely to due to $\text{PM}_{2.5}$ from HAP in men as well as women. However, we lack personal $\text{PM}_{2.5}$ measurements from HAP among men, as they generally do not cook and most data comes from women. Men would be expected to be exposed to less $\text{PM}_{2.5}$ due to biomass burning than women, but still have appreciable exposures well above WHO recommended levels.

3.2.2. Birthweight

The combined data from Wylie et al. (2017), Sun et al. (2016), and Jedrychowski et al. (2004), indicate a median decrease in birth weight of those exposed to 70 $\mu\text{g}/\text{m}^3$ (typical of LPG stoves), vs. 270 $\mu\text{g}/\text{m}^3$ (typical of unimproved stoves), of 343 g (2.5th% –344, 97.5th% 1114). Again, however the simulation interval includes the null value of 0.

Reduced birthweight is very strongly associated with high infant mortality, although the mechanisms are not well understood, and the association might not be causal (Wilcox, 2001; Hernández-Díaz et al., 2008). Reduced birthweight has also been associated with disease later in life (Gluckman et al., 2008). There is also some evidence of a direct effect of air pollution on infant mortality (Romieu et al., 2012; Bruce et al., 2013), although the data are sparse and it is not clear if that effect might be mediated by maternal factors and reduced birth weight. A direct causal effect of birthweight on infant mortality is made more plausible assuming some of those with reduced birthweight are presumably pre-term birth, given that there is evidence that both ambient $\text{PM}_{2.5}$ (Malley et al., 2017) and HAP (Amegah et al., 2014), are associated with pre-term birth, and pre-term birth is strongly related to

Table 1

Simulation results: expected outcomes at different levels of exposure for physician-diagnosed childhood pneumonia, systolic blood pressure, and birthweight.

$\text{PM}_{2.5}$ $\mu\text{g}/\text{m}^3$	$\log \text{PM}_{2.5}$	RR childhood pneumonia*, median, 2.5th and 97.5th percentile	SBP (mm) women > 50**, median, 2.5th and 97.5th percentile	Birthweight (grams)***, median, 2.5th and 97.5th percentile
270	5.6	2.81 (1.23–3.91)	131.4 (120.4–142.6)	2543 (1874, 3102)
120	4.8	2.24 (1.06–3.51)	128.1 (117.9–137.9)	2770 (2306,3206)
70	4.2	1.78 (1.05–3.14)	125.8 (115.9–134.7)	2880 (2507, 3264)
35	3.6	1.41 (1.01–2.24)	123.3 (113.2–130.4)	3003 (2722, 3367)
10	2.3	1.02 (1–1.07)	118	3200

*Burnett et al., 2014; reference group was exposed at 6 $\mu\text{g}/\text{m}^3$.

**From Baumgartner et al., 2011, reference level of 118 mm Hg taken from lower confidence level of mean SBP.

***From Wylie et al., 2017, Sun et al., 2016, Jedrychowski et al., 2004, reference weight of 3200 g taken from mean birthweight in Wylie et al. (2017).

infant mortality. WHO estimates that about 12% of births in developing countries are pre-term and estimates that about 50% of infant mortality worldwide is due to pre-term birth (WHO, 2012).

If we assume a causal effect of low birth weight on infant mortality, given that there is a drop of one unit (per 1000) in infant mortality for each 30 g of birthweight (Wilcox, 2001), an improvement of 343 g (LPG stoves vs traditional stoves, assuming the point estimate is accurate) would translate into an approximate drop in infant mortality of 11/1000 live births. To put this in perspective, the worldwide infant mortality rate in 2015 was 32/1000 (World Bank, 2017).

A more conservative estimate of effect on infant mortality might consider that among low birthweight babies, only pre-term births affect infant mortality. Sun et al. (2015) have estimated the exposure-response curve for PM_{2.5} and pre-term birth in a meta-analysis of 13 ambient air studies. They found a 1.13 increased odds of a pre-term birth for every 10 µg/m³ increment of PM_{2.5}. While this assumption of a linear increase in log odds may hold true for the relatively low levels of PM_{2.5} in ambient air, it may not hold for higher levels of PM_{2.5}, as noted by Malley et al. (2017), given the high exposure-response slopes at low levels followed by a plateau seen for the odds ratio for a variety of outcomes in relation to PM_{2.5}. We have estimated a modified exposure-response curve from Sun et al. (2015) using log-transformed PM_{2.5}, such that it crosses the original curve from Sun et al. at the 50 µg/m³ point, and then starts to trail off (an odds ratio of 2.00 for each log 10 units of PM_{2.5}) (see supplemental Fig. 1).

Using this modified exposure-response curve, and assuming a baseline rate of 7% pre-term birth at 10 µg/m³ based on European countries with low pre-term birth rates (WHO, 2017), we find that an exposure of 270 µg/m³ (from traditional stoves) would be estimated to have a pre-term birth rate of 18.8%, compared to an estimated rate of 12.6% at 70 µg/m³ (a change of 6.2%). Assuming that 7.3% of pre-term births die as infants (WHO, 2012), out of 1000 live births, we would expect 13.7 (188*0.073) deaths among those exposed at 270 µg/m³, versus 9.2 deaths among those exposed at 70 µg/m³, an estimated decrease in 4.5 deaths per 1000 live births.

3.2.3. Pneumonia

From the point of view of subsequent mortality, severe infant pneumonia cases are of primary interest, compared to all pneumonia. Severe infant cases are the most likely to die, with an estimated 9% mortality rate (Walker et al., 2013). The baseline severe infant pneumonia rate (physician diagnosed) was 260 cases per 1000 child years among the control group in the 2002–2004 RESPIRE study in Guatemala. However, in recent years the pneumococcal polysaccharide vaccine (PPV) has been introduced, lowering pneumonia rates (Berical et al., 2016). More recent data in rural areas of other low/middle income countries suggest that a current baseline rate of severe pneumonia in the first 2 years of life is likely to be closer to 50 to 60 cases per 1000 children (Faroqui et al., 2015; Mackenzie et al., 2014).

We have roughly assumed that those with the highest personal exposure (300 µg/m³) have a rate of 80 cases per 1000, and then used the relative risks from Burnett et al. (2014) to estimate the rate for lower levels of exposure. Table 2 shows the number of excess cases per 1000 child years with exposures above 10 µg/m³, using the rate ratios from Table 1.

From Table 2 we see that the estimated decrease in cases from an exposure of 270 µg/m³ (estimated exposure for traditional stove users) to an exposure of 70 µg/m³ (estimated exposure for LPG users) is 29 cases of severe pneumonia. Given the estimated 9% mortality rate (Walker et al., 2013), an LPG intervention would be expected to avoid about 2–3 deaths from infant pneumonia among per year among 1000 children.

4. Discussion

This study demonstrates the hypothetical impacts of LPG use on

Table 2

Pneumonia cases expected per year among 1000 children, and excess cases compared to a level of 10 µg/m³ PM_{2.5}^a.

PM _{2.5}	RR ^b	Severe cases/yr
300	2.88	80
270	2.81 (1.23–3.91)	78 (34–109)
120	2.24 (1.06–3.51)	62 (29–98)
70	1.78 (1.05–3.14)	49 (29–87)
35	1.41 (1.01–2.24)	39 (28–62)
10 ^a	1.02 (1–1.07)	28 (28–30)

^a 10 µg/m³ is the WHO recommended level.

^b RRs from simulations, as per Table 1.

exposure and health. LPG use is increasing worldwide as households move up the energy ladder. The feasibility of LPG being adopted, in rural areas now using biomass, is increasing as gas prices are lowered. However, LPG still is more expensive than biomass in much of the world, making price an issue in adaptation. Countries like India and Peru have begun nationwide subsidization of LPG, which should have substantial impact.

We argue that use of LPG would have important health benefits compared to biomass stoves even if PM_{2.5} levels are not lowered to WHO standards, due to stove stacking and other exposure via ambient air.

Despite the uncertainty around the relationship between HAP exposure and health, and in accordance with others (Bruce et al., 2015, Johnson and Chiang, 2015), our results provide supportive evidence that appreciable health gains can be achieved at reduced PM_{2.5} concentrations that are still higher than WHO recommend interim levels of exposure (35 µg/m³). At 70 µg/m³, using our best estimate of the expected personal PM_{2.5} exposure in homes with LPG stoves (based only a single study), our simulations indicate that there is an expected reduction of approximately 37% of infant pneumonia cases (avoiding 2–3 deaths annually per 1000 infants), an expected 5.5 mm Hg reduction in SBP among older women, and an expected 338 g increase in birthweight, compared to traditional unimproved stoves (270 µg/m³). Improvements in these three outcomes can be measured in the short-term in an intervention study. The reductions in SBP and the increases in birthweight would be expected to lead to later decreases in morbidity and mortality.

We have noted that the evidence suggests that age is a potential modifier of the expected blood pressure benefits to be achieved following a clean fuel cookstove intervention, with benefits accruing primarily to older women. Evidence from the ambient air pollution field suggests that other factors may also confer increased susceptibility to the adverse health effects of higher air pollution exposures (Clark and Peel, 2014). Large robust interventions, which permit sub-group analyses, are needed to understand if these same susceptibility factors may also lead to larger health benefits following reductions in household air pollution. Potential modifying factors include poor nutrition, psychosocial stressors, other environmental pollutants, and comorbid conditions, all of which have been implicated in the adverse outcomes typically associated with HAP (Clark and Peel, 2014).

There are two important sources of uncertainty in our estimates, which led us to use simulations to estimate health effects in order to incorporate both of them. One is the estimate of the exposure level resulting in our estimates of personal exposure levels from traditional biomass stoves and from LPG stoves, the latter based on only one study. We reflect this uncertainty by assuming that our estimates are drawn from normal distributions in which the standard deviation equals the mean, which reflects high uncertainty, and also reflects what has been observed in exposure studies. The high observed standard deviations of personal PM_{2.5} measurements in any given study stem from a number of factors, one of the most important being that only one or a small number of measurements are collected per person, so that between-

person variability in a given study may reflect a lot of within-person-variability. For example, if measurements are collected over a period of time in different seasons, we know there will be marked variability due to season, with higher PM_{2.5} being present in the winter (Ni et al., 2016; Huang et al., 2017; Carter et al., 2016), and many subjects may have measurement from only one season. Similarly, cooking practices may differ one day to the next, and sparse measurements will reflect such variability. Compliance may also vary person-to-person. Besides such within-person variation within a given study, there is a very wide range of results across studies (Balakrishnan et al., 2014; Rehfuess et al., 2014). This is also likely to reflect these same sources of variation – e.g., if one study was conducted only in winter and another was conducted in summer.

The other important source of uncertainty in our results is that our exposure-response data are based on a small number of studies. Sample sizes for these studies were reasonable and confidence intervals relatively tight, but nonetheless confidence in results is limited by the small number of studies. For SBP, we used an exposure-response curve based on only a single study, albeit with good supporting evidence from other studies. For childhood pneumonia, the exposure-response data are largely from a single HAP study, supplemented by other studies of infant pneumonia in relation to ambient air and second-hand tobacco smoke. For birthweight, the exposure-response data are based on three studies. This reliance on only a few studies is due to the paucity of studies with PM_{2.5} personal exposure data used in exposure-response analyses. Consistency of effect needs to be demonstrated in other studies of HAP and additional settings, before strong inferences can be made that the associations observed are causal.

Of the two sources of uncertainty in our simulations, the uncertainty in exposure is much greater than the uncertainty due to exposure-response coefficients. While we have focused on the point estimates in the expected amount of improvement in health outcomes, our expected improvements include the null value of 0 for all outcomes. For example, the expected improvement in SBP due to lowering PM_{2.5} levels from 270 µg/m³ to 70 µg/m³ is 5.5 mm Hg, with a 90% simulation interval of –6.2 to 18.0 mm Hg. Without the uncertainty in exposure, the corresponding expected improvement is again 5.5 mg Hg, but the 90% simulation is 2.5 to 8.4 mg Hg, excluding the null. The width of this interval increases 4-fold after incorporating uncertainty regarding exposure.

Of note is that our estimate of the improvement of birth weight (338 g increase) due to reduced exposure to PM_{2.5} is greater than the increase in birthweight estimated due to mothers stopping smoking (200 g, US Surgeon General, 1990), and 4–5 times greater than the effect due to exposure to second-hand smoke (60 g, Salmasi et al., 2010). The increase in weight due to decreasing biomass PM_{2.5} exposure seems disproportionate as compared to the decrease due to stopping maternal smoking, given the presumed likely amounts of exposure to the fetus. However, the components of tobacco smoke vs. biomass burning may affect the fetus differently. There is some evidence from animal studies, for example, that biomass causes a stronger inflammatory response in the lung than cigarette smoke (Mehra et al., 2012); similar seemingly disproportionate effects might be seen for the fetus.

We have noted the high variability in estimates of personal exposure in any given study. In attempting to assess a long term average, which may be the exposure of interest for many outcomes, measurement error is inevitable in due to the few samples obtained from individuals. For some outcomes (e.g., birth outcomes) we do not know the proper time window for the most relevant exposure. Such error in the measurement of personal exposure is likely to be of the classical type, would be expected to be non-differential, and would therefore be expected to bias exposure-response coefficients to the null. Beside measurement error for exposure, there may be error in measuring outcome; for example ascertainment of severe infant pneumonia is not straightforward.

Furthermore, PM_{2.5} may in fact not be the most appropriate

summary measure of HAP's impact on health. We need further exposure-response data in HAP settings which include assessment of other contaminants (e.g., PAHs, NO₂, black carbon, CO) (Naehler et al., 2007). Even if PM_{2.5} is the best measure of HAP exposure (most predictive in relation to health), it would be useful to have PM_{2.5} analyses that include data on composition and source apportionment; composition and sources of PM_{2.5} vary by setting, and thus toxicity, mutagenicity, and oxidative potential may also vary.

Other sources of uncertainty include possible biases in observational studies due to confounding or selection bias. One cause of concern in relation to observational studies is possible bias due to socio-economic status, which may be associated with clean fuel use and lower PM_{2.5} exposure. Regarding the studies used here, however, confounding by SES seems relatively unlikely. The studies used here adjusted for SES in their analyses, often via education, although residual confounding is still possible. The studies used here for the most part included populations which were quite homogenous, e.g., all with biomass stoves (i.e., comparing effects across a range of biomass-related PM_{2.5} concentrations), or intervention studies where the improved or clean fuel stove was disseminated among biomass stove users; this homogeneity is another factor likely to limit the type of confounding often present in studies observationally comparing users of different stove types or fuels.

In conclusion, despite all the uncertainties noted above that are inherent in the data, existing evidence indicates that substantial exposure reductions and health gains can be expected from an LPG intervention.

Conflict of interest

The authors have no conflict of interest.

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